STAT475/675 TUT09

http://www.sfu.ca/~zza115/teaching.html http://people.stat.sfu.ca/~joanh/stat475-675web.html

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Three-Way Contingency Table with Poisson-distributed Cell Counts

- Cell count N with marginal X, Y and Z
- $\mu_{ijk} = E(N_{ijk}) = E(N|X=i, Y=j, Z=k)$
- Assume $N_{ij} \sim \text{Poisson}(\mu_{ijk})$
- Saturated Loglinear Model (XYZ):

$$\ln \mu_{ijk} = \beta_0 + \beta_i^X + \beta_j^Y + \beta_k^Z + \beta_{ij}^{XY} + \beta_{jk}^{YZ} + \beta_{ik}^{XZ} + \beta_{ijk}^{XYZ}$$

- Number of non-redundant parameters: IJK
- Mutual independence of X, Y and Z: LRT and χ^2 test
- Loglinear Model of Homogeneous Association (XY, YZ, XZ)

$$\ln \mu_{ijk} = \beta_0 + \beta_i^X + \beta_j^Y + \beta_k^Z + \beta_{ij}^{XY} + \beta_{jk}^{YZ} + \beta_{ik}^{XZ}$$

- Number of non-redundant parameters: IJK (I-1)(J-1)(K-1)
- When I = J = 2,

$$\ln \theta_{XY(k)} = \ln \frac{\mu_{11k}\mu_{22k}}{\mu_{12k}\mu_{21k}} = \beta_{11}^{XY} + \beta_{22}^{XY} - \beta_{12}^{XY} - \beta_{21}^{XY}$$

- * $\theta_{XY(k)}$ stays still for all k, i.e., homogeneous conditional OR holds * if $\beta_{ij}^{XY} = 0$ for all i, j, then $\cdot \ \theta_{XY(k)} = 1 \text{ for all } k$ $\cdot \ X \perp\!\!\!\perp Y|Z$
- - · consider model (YZ, XZ)
- Loglinear Model of Independence (X, Y, Z)

$$\ln \mu_{ijk} = \beta_0 + \beta_i^X + \beta_j^Y + \beta_k^Z$$

- Number of non-redundant parameters: 1 + (I-1) + (J-1) + (K-1) = I + J + K - 2

Demo I

Data "UCBAdmissions" (included in R default Package "datasets") is on applicants to graduate school at Berkeley for the six largest departments in 1973 classified by admission and sex.

- Admit: Admitted, Rejected
- Gender: Male, Female
- Dept: A, B, C, D, E, F

Poisson Rate Regression

• $Y|t, x_1, \ldots, x_p \sim \text{Poisson}(\mu(t, x_1, \ldots, x_p))$ and assume

$$\ln \mu(t, x_1, \dots, x_p) = \ln t + \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p$$

or equivalently,

$$E(Y/t|t, x_1, ..., x_p) = \frac{\mu(t, x_1, ..., x_p)}{t} = \exp(\beta_0 + \beta_1 x_1 + ... + \beta_p x_p)$$

or

$$E(Y|t, x_1, ..., x_p) = \mu(t, x_1, ..., x_p) = t \exp(\beta_0 + \beta_1 x_1 + ... + \beta_p x_p)$$

- Y: count of events
- -t: measure of opportunity for events

Demo II

250 groups went to a park. Each group was questioned about

- count (integer): number of fishes they caught;
- persons (integer): number of people were in the group;
- camper (categorical): whether or not they brought a camper;
- livebait (categorical): whether or not they used live bait;
- child (categorical): number of children were in the group.

See https://stats.idre.ucla.edu/r/dae/zip/ for more details.

To-do:

- Predict the number of fish caught by loglinear models.
- Take variable "persons" as t and built a Poisson rate model.

Two Problems with Poisson Regression

- Overdispersion: the equality of mean and variance is violated
 - solutions, e.g.
 - * negative binomial distribution
 - * quasi-likelihood estimation
- Zero inflation: too many zeros are observed in response
 - solutions, e.g.
 - * zero-inflated Poisson (ZIP) model (Lambert, 1992): a mixture distribution of the form

$$Y \begin{cases} = 0 \text{ with probabity } \pi \\ \sim \text{Poisson}(\mu) \text{ with probabity } 1 - \pi \end{cases}$$

or equivalently,

$$\Pr(Y = y) = \begin{cases} \pi + (1 - \pi) \exp(-\mu), \ y = 0\\ \frac{(1 - \pi)\mu^y \exp(-\mu)}{y!}, \ y \in \mathbb{N} \end{cases}$$

where
$$\pi = \pi(z_1, ..., z_J) = \text{logit}^{-1}(\gamma_0 + \gamma_1 z_1 + ... + \gamma_J z_J)$$
 and $\mu = \mu(x_1, ..., x_p) = \exp(\beta_0 + \beta_1 x_1 + ... + \beta_p x_p)$

Demo II (Continued)

To-do:

• Consider a ZIP model and predict again the probability that a group caught zero fish.

Generalized Linear Models

- Special cases
 - Ordinal linear model
 - Logit model
 - Baseline-category logit Model
 - Cumulative logit Model
 - Adjacent-categories logit Model
 - Loglinear model
- Unified framework:

$$g_i(\mu_1, \dots, \mu_I) = \beta_0 + \beta_{i1}x_1 + \dots + \beta_{ip}x_p, \quad j = 1, \dots, J,$$

with $Y \sim f_{\mu_1,\dots,\mu_I}$, a parametric distribution (belonging to the exponential family) characterized by with $I \hookrightarrow J_{\mu_1,\dots,\mu_I}$, a parametric distribution (belonging to the exponential rainity) character $\mu_i = \mu_i(x_1,\dots,x_p), \quad i=1,\dots,I$ - random component: $Y \sim f_{\mu_1,\dots,\mu_I}$ - systematic component: $\beta_0 + \beta_{j1}x_1 + \dots + \beta_{jp}x_p$ - link function (usually monotone and differentiable over the range of (μ_1,\dots,μ_I)): $g_j(\cdot)$